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## 6.1 The Probability of an Event

**Definition 6.1.** An outcome in a finite sample space may be assigned a **probability mass**, a real number in [0, 1] such that the sum of the probabilities of all outcomes in S is 1. The **probability** of an event A is the sum of the probability masses of all outcomes in A, subject to

$$0 \le P(A) \le 1 \quad P(\emptyset) = 0 \quad P(S) = 1.$$

If  $A_1, A_2, \ldots$  is a sequence of mutually exclusive events, then

$$P(A_1 \cup A_2 \cup \cdots) = P(A_1) + P(A_2) + \cdots$$

If an experiment can result in N equally likely outcomes, the probability of an event A which contains n outcomes is

$$P(A) = \frac{n}{N}$$

We have the following collection of results.

**Theorem 6.2.** For any two events A and B,  $P(A \cup B) = P(A) + P(B) - P(A \cap B)$ .

This can be proved in a number of ways, for example by use of a Venn Diagram.

**Corollary 6.3.** If A and B are mutually exclusive,  $P(A \cup B) = P(A) + P(B)$ .

**Corollary 6.4.** If  $A_1, A_2, ..., A_n$  are mutually exclusive,  $P(A_1 \cup A_2 \cup \cdots \cup A_n) = P(A_1) + P(A_2) + \cdots + P(A_n)$ . **Theorem 6.5.**  $P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(A \cap B) - P(A \cap C) - P(B \cap C) + P(A \cap B \cap C)$ .

This can be proved via Venn diagrams, although a generalized Inclusion-Exclusion rule can be proved by induction. A set-based proof follows from some set operations. See if you can justify each of the steps; you need to know how distribution of intersection and union works:  $A \cup (B \cap C) = (A \cup B) \cap (A \cup C)$  and  $A \cap (B \cup C) = (A \cap B) \cup (A \cap C)$ .

$$\begin{aligned} P(A \cup B \cup C) &= P(A \cup (B \cup C)) \\ &= P(A) + P(B \cup C) - P(A \cap (B \cup C)) \\ &= P(A) + P(B) + P(C) - P(B \cap C) - P((A \cap B) \cup (A \cap C))) \\ &= P(A) + P(B) + P(C) - P(B \cap C) - (P(A \cap B) + P(A \cap C) - P(A \cap B) \cap (A \cap C))) \\ &= P(A) + P(B) + P(C) - P(B \cap C) - P(A \cap B) - P(A \cap C) + P(A \cap B \cap C)) \end{aligned}$$

**Theorem 6.6.** Complement Rule: P(A) + P(A') = 1, *i.e.* P(A) = 1 - P(A').

This follows from the fact that  $A \cup A' = S$  and P(S) = 1.

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## 6.2 Conditional Probability and Independence

**Definition 6.7.** The probability that event A occurs when it is known that event B has occurred is a conditional probability and it is written P(A|B), read "The probability of A given B".

**Theorem 6.8.** The conditional probability P(A|B), assuming P(B) > 0, may be calculated as

$$P(A|B) = \frac{P(A \cap B)}{P(B)}.$$

This can be proved in the following way. Because we know that B has occurred, we restrict our sample space to event B. Then event A occurring given B means that both A and B occur.

**Definition 6.9.** Two events *A* and *B* are said to be **independent** if the probability of one event occurring is not affected by whether or not the other event occurs, or in notation

$$P(A|B) = P(A)$$
 and  $P(B|A) = P(B)$ .

Otherwise, the events are **dependent**.

WARNING: Be careful not to confuse "mutually exclusive" and "independent!" Not only are they **not** synonymous, two events can **never** be both independent and mutually exclusive, unless one of the events has probability zero.

This leads to a corollary.

Corollary 6.10. Two events A and B are independent if and only if

$$P(A)P(B) = P(A \cap B).$$

You can derive this by using the formulaic definition of conditional probability and replacing P(A|B) with P(A).

And the formula for conditional probability may be re-written giving us another corollary.

**Corollary 6.11.** The Product Rule: For any events A and B,  $P(A \cap B) = P(A|B)P(B) = P(B|A)P(A)$ .

This can be extended using induction

**Theorem 6.12.** For events  $A_1, A_2, \ldots, A_k$ ,

$$P(A_1 \cap A_2 \cap \dots \cap A_k) = P(A_1)P(A_2|A_1)P(A_3|A_1 \cap A_2) \cdots P(A_k|A_1 \cap A_2 \cap \dots \cap A_{k-1}).$$

If the events are independent then

$$P(A_1 \cap A_2 \cap \dots \cap A_k) = P(A_1)P(A_2) \cdots P(A_k)$$

(but be careful, the converse is not true).

**Definition 6.13.** The events  $\mathcal{A} = \{A_1, A_2, \dots, A_k\}$  are **mutually independent** if and only if for any subset  $\{A_{i_1}, \dots, A_{i_r}\}$ ,

$$P(A_{i1} \cap \cdots \cap A_{i_r}) = P(A_{i_1}) \cdots P(A_{i_r}).$$

## 6.3 Bayes Rule

**Definition 6.14.** A partition of a set S is a collection of subsets  $E_1, E_2, \ldots E_k$  such that any two distinct  $E_i$  and  $E_j$  are mutually exclusive (disjoint) and  $E_1 \cup E_2 \cup \cdots \cup E_k = S$ .

**Theorem 6.15.** The Law of Total Probability: Let  $B_1, \ldots, B_k$  partition S. Then for any event A,

$$P(A) = \sum_{i=1}^{k} P(A \cap B_i) = \sum_{i=1}^{k} P(B_i) P(A|B_i)$$

This can easily be seen to be true since  $A \cap B_i$  and  $A \cap B_j$  are disjoint for  $i \neq j$ :

$$(A \cap B_i) \cap (A \cap B_j) = A \cap (B_i \cap B_j) = A \cap \emptyset = \emptyset,$$

and the rest follows from straightforward application of additivity theorems.

**Theorem 6.16.** Bayes Rule: Let  $B_1, \ldots, B_k$  partition S. Then for any event A,

$$P(B_r|A) = \frac{P(B_r \cap A)}{\sum_{i=1}^k P(B_i \cap A)} = \frac{P(B_r)P(A|B_r)}{\sum_{i=1}^k P(B_i)P(A|B_i)}.$$

This is proved by applying the law of total probability to the formula for conditional probability. It can be intuitively justified with a tree diagram, although this leads to some counter-intuitive examples.